

Gibbs point processes : modelling and inference

J.-F. Coeurjolly (Grenoble University)

et J.-M Billiot, D. Dereudre, R. Drouilhet, F. Lavancier

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Outline

- 1 Type of data of interest
- 2 Gibbs models
 - Brief background
 - Examples
- 3 Identification
 - Maximum likelihood
 - Pseudo-likelihood method
 - Takacs-Fiksel method
 - Variational Principle method
- 4 Validations through residuals
 - Residuals for spatial point processes
 - Measures of departures to the true model
 - Asymptotics

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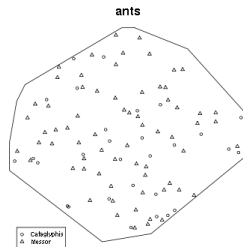
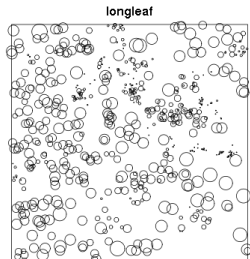
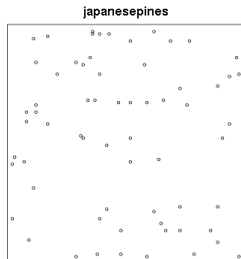
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Type of data

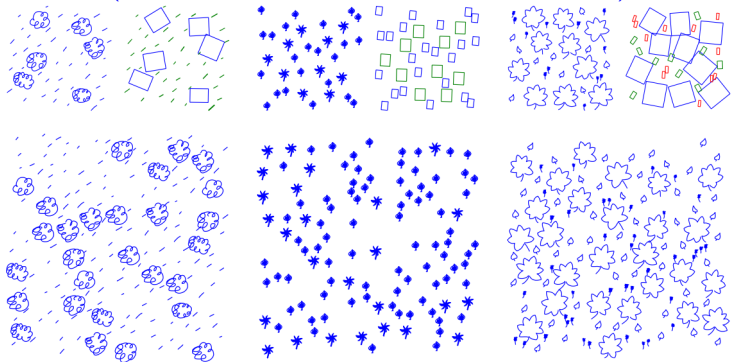


Scientific questions :

- 1 Independence or interaction ? regular or clustered distribution ?
- 2 Spatial variation in the density and marks ?
- 3 Interaction in each sub-pattern, between sub-patterns...

Another example in computer graphics

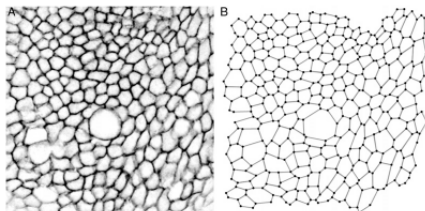
(Hurtut, Landes, Thollot, Gousseau, Drouilhet, C.'09.)



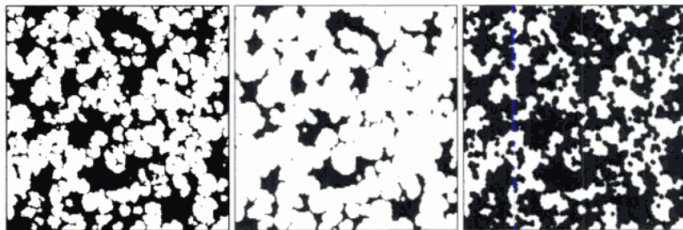
Problem : simulation at large scale of a pattern drawn in a small window.

Examples with geometrical structures

① Epithelial cells



② Materials interface



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Marked point processes

- State space : $\mathbb{S} = \mathbb{R}^d \times \mathbb{M}$ associated to $\mu = \lambda \otimes \lambda^m$.
- Let $x^m = (x, m)$ an element of \mathbb{S} , i.e. a marked point.
- Ω is the space of locally finite point configurations φ in \mathbb{S}
- For $\Lambda \in \mathbb{R}^d$: Λ bounded borelian set of \mathbb{R}^d .
- φ_Λ is the restriction of φ on Λ and $|\varphi_\Lambda|$ is the number of points φ_Λ .

Definition (marked point process)

A marked point process is a random variable, Φ , on Ω .

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Poisson process

For $z > 0$, the standard (non-marked) poisson point process π^z with intensity $z\lambda$ is defined by

$$\begin{cases} \forall \Lambda, & |\pi_\Lambda^z| := \pi^z(\Lambda) \sim \mathcal{P}(z\lambda(\Lambda)) \\ \forall \Lambda, \Lambda' \text{ with } \Lambda \cap \Lambda' = \emptyset, & \pi_\Lambda^z \text{ and } \pi_{\Lambda'}^z \text{ are independent.} \end{cases}$$

Stationary Gibbs models on \mathbb{R}^d

Consider a parametric family of energies $(V_\Lambda(\cdot; \theta))_{\Lambda \in \mathbb{R}^d}$ for $\theta \in \mathbb{R}^p$ defined on Ω and with value on $\mathbb{R} \cup \{+\infty\}$.

Definition

A probability measure P_θ on Ω is a stationary marked Gibbs measure for the compatible and invariant by translation family of energies $(V_\Lambda(\cdot; \theta))_{\Lambda \in \mathbb{R}^d}$ if for every $\Lambda \in \mathbb{R}^d$, for P_θ -a.e. outside configuration φ_{Λ^c} , the distribution of P_θ conditionally to φ_{Λ^c} admits the following conditional density with respect to π_Λ ($:= \pi_\Lambda^1$):

$$f_\Lambda(\varphi_\Lambda | \varphi_{\Lambda^c}; \theta) = \frac{1}{Z_\Lambda(\varphi_{\Lambda^c}; \theta)} e^{-V_\Lambda(\varphi; \theta)},$$

where $Z_\Lambda(\varphi_{\Lambda^c}; \theta)$ is the normalizing constant called partition function.

Compatibility for all $\Lambda \subset \Lambda' \in \mathbb{R}^d$, there exists a measurable function $\psi_{\Lambda, \Lambda'}$ from Ω to $\mathbb{R} \cup \{+\infty\}$ such that

$$\forall \varphi \in \Omega, \quad V_{\Lambda'}(\varphi; \theta) = V_\Lambda(\varphi; \theta) + \psi_{\Lambda, \Lambda'}(\varphi_{\Lambda^c}; \theta).$$

Existence/uniqueness problem

The choice of $(V_\Lambda(\cdot; \theta))_{\Lambda \in \mathbb{R}^d}$ entirely defines the Gibbs measure P_θ .

But, given a family $\{V_\Lambda(\cdot; \theta)\}_{\Lambda \in \mathbb{R}^d}$: does there exist a Gibbs measure P_θ ? Is it unique (phase transition problem)?

- [Ruelle\(69\)](#), [Preston\(76\)](#) : Superstable and lower regular potentials (e.g. Lennard-Jones model).
- [BBD\(99\)](#) : local stability and finite range (actually quasilocality) assumptions.
- [D\(05\)](#) : non hereditary energies. [A family $(V_\Lambda(\cdot; \theta))_{\Lambda \in \mathbb{R}^d}$ is **hereditary** if $\forall \Lambda \in \mathbb{R}^d, V_\Lambda(\varphi; \theta) = +\infty \Rightarrow V_\Lambda(\varphi \cup x^m; \theta) = +\infty.$]
- [DDG\(10\)](#) : stability and locality (to the configuration) assumptions.

Existence assumption [Mod]

we observe a realization of Φ with marked Gibbs measure P_{θ^*} , where $\theta^* \in \mathring{\Theta}$, Θ is a compact set of \mathbb{R}^p and for all $\theta \in \Theta$, there exists a stationary marked Gibbs measure P_θ for the family $(V_\Lambda(\cdot; \theta))_{\Lambda \in \mathbb{R}^d}$ assumed to be invariant by translation, compatible.

Local energy function

Definition

The energy required to insert x^m in φ is defined for every $\Lambda \ni x$ by

$$V(x^m|\varphi; \theta) := V_\Lambda(\varphi \cup x^m; \theta) - V_\Lambda(\varphi; \theta),$$

which from the compatibility assumption is independent of Λ .

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For every mark m , every configuration φ and for all $\theta \in \Theta$

[LS] Local stability : $\exists K > 0$ such that $V(0^m|\varphi; \theta) \geq -K$.

[FR] Finite range : $\exists D > 0$ such that $V(0^m|\varphi; \theta) = V(0^m|\varphi_{\mathcal{B}(0,D)}; \theta)$.

D, K independent of θ, m, φ .

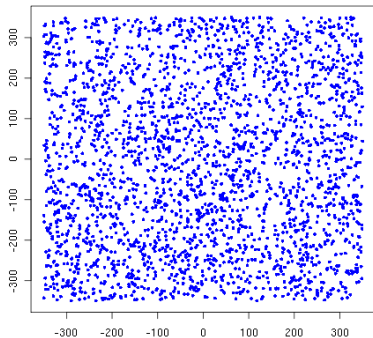
BBD(99)

$$\mathbf{[LS]} + \mathbf{[FR]} \implies \mathbf{[Mod]}$$

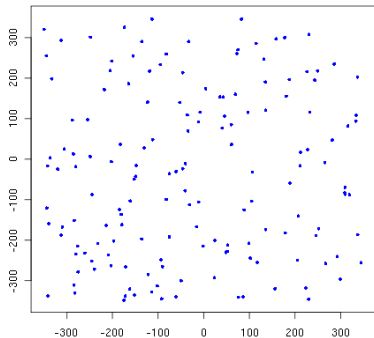
Poisson point process, $\mathbb{M} = \{0\}$

$$V_{\Lambda}(\varphi; \theta) = \theta_1 |\varphi_{\Lambda}|$$

$\theta_1=5$



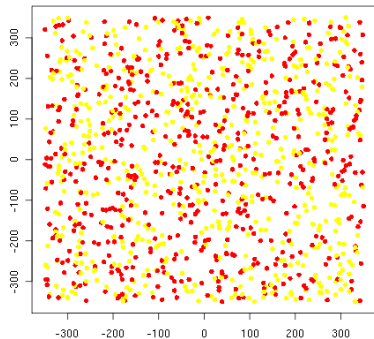
$\theta_1=8$



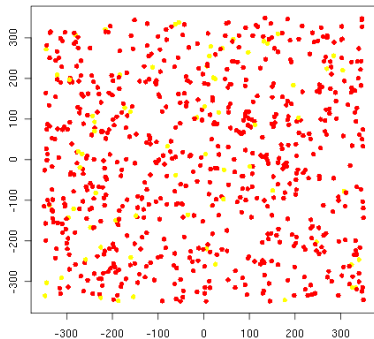
Multi-type Poisson point process, $\mathbb{M} = \{1, 2\}$

$$V_{\Lambda}(\varphi; \theta) = \theta_1^1 |\varphi_{\Lambda}^1| + \theta_1^2 |\varphi_{\Lambda}^2|$$

$$\theta_1^1=6, \theta_1^2=6$$



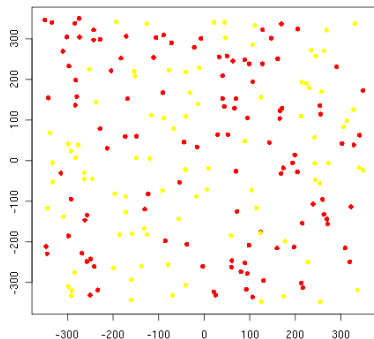
$$\theta_1^1=6, \theta_1^2=8$$



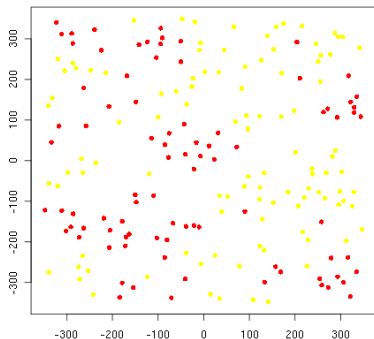
Strauss marked point process, $\mathbb{M} = \{1, 2\}$

$$V_\Lambda(\varphi; \theta) = \sum_{m=1}^2 \theta_1^m |\varphi_\Lambda^m| + \sum_{1 \leq m \leq m' \leq 2} \theta_2^{m,m'} \sum_{\substack{\{x^m, y^{m'}\} \in \mathcal{P}_2(\varphi) \\ \{x^m, y^{m'}\} \cap \Lambda \neq \emptyset}} \mathbf{1}_{[0, D^{m,m'}]}(\|y - x\|),$$

$$\theta_1 = (2, 2), \theta_2 = (2, 2, 2)$$



$$\theta_1 = (2, 2), \theta_2 = (6, 6, 10)$$

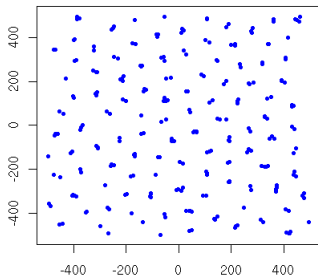


Multi-Strauss point process, $\mathbb{M} = \{0\}$ on a planar structured graph

$$V_{\Lambda}(\varphi; \theta) = \theta_1 |\varphi_{\Lambda}| + \sum_{k=1}^K \theta_2^k \sum_{\substack{\{x,y\} \in G_2(\varphi) \\ \{x,y\} \cap \Lambda \neq \emptyset}} \mathbf{1}_{[D_{k-1}, D_k]}(\|y - x\|),$$

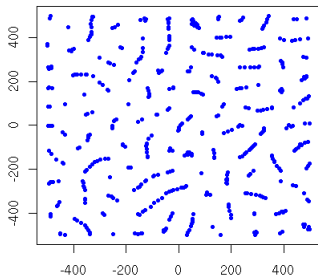
$$G_2(\varphi) = \mathcal{P}_2(\varphi)$$

number of points=266



$$G_2(\varphi) = Del_2(\varphi)$$

number of points=472

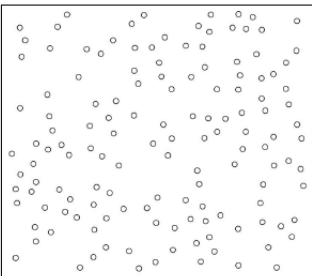


$$D_0 = 0, D_1 = 20, D_2 = 80, \theta = (1, 2, 4)$$

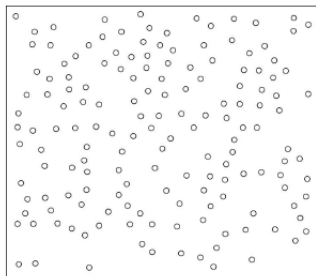
Lennard-Jones model

$$V_{\Lambda}^{LJ}(\varphi; \theta) := \theta_1 |\varphi_{\Lambda}| + 4\theta_2 \sum_{\substack{\{x,y\} \in \mathcal{P}_2(\varphi) \\ \{x,y\} \cap \Lambda \neq \emptyset}} \left(\left(\frac{\theta_3}{\|y-x\|} \right)^{12} - \left(\frac{\theta_3}{\|y-x\|} \right)^6 \right)$$

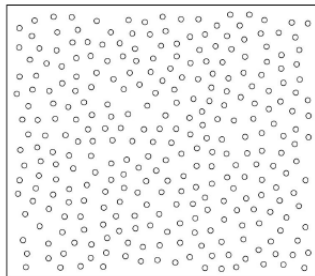
with $\theta = (\theta_1, \theta_2, \theta_3) \in \mathbb{R} \times (\mathbb{R}^+)^2$.



$\theta_2 = 0$



$\theta_2 = 0.1$

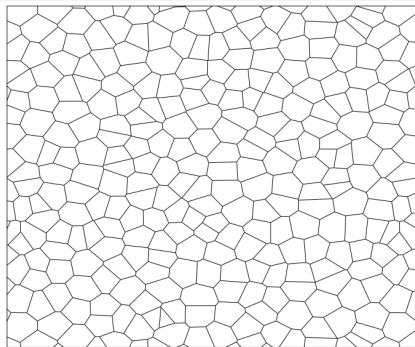


$\theta_2 = 2$

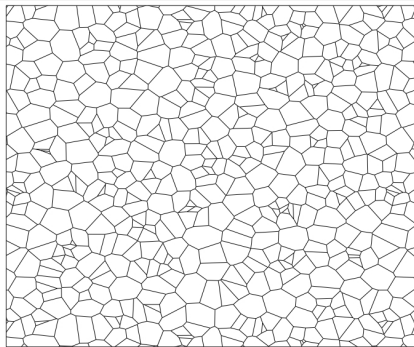
Gibbs Voronoi tessellation

$$V_{\Lambda}(\varphi) = \sum_{\substack{C \in \text{Vor}(\varphi) \\ C \cap \Lambda \neq \emptyset}} V_1(C) + \sum_{\substack{C, C' \in \text{Vor}(\varphi) \\ C \text{ and } C' \text{ are neighbors} \\ (C \cup C') \cap \Lambda \neq \emptyset}} V_2(C, C')$$

$V_1(C)$: deals with the shape of the cell and $V_2(C, C') = \theta d(\text{vol}(C), \text{vol}(C'))$.



$\theta > 0$



$\theta < 0$

Quermass model $\mathbb{M} = [0, \bar{R}]$

For finite configuration φ

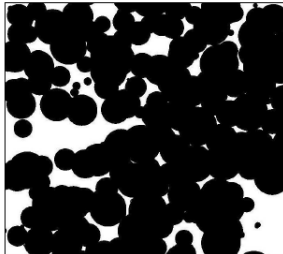
$$V(\varphi; \theta) = \theta_1 |\varphi| + \theta_2 \mathcal{P}(\Gamma) + \theta_3 \mathcal{A}(\Gamma) + \theta_4 \mathcal{E}(\Gamma) \quad \text{where } \Gamma = \bigcup_{(x,R) \in \varphi} \mathcal{B}(x, R)$$

where $\mathcal{P}(\Gamma)$, $\mathcal{A}(\Gamma)$ and $\mathcal{E}(\Gamma)$ respectively denote the perimeter, the volume and the Euler-Poincaré characteristic of Γ .

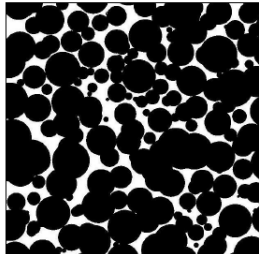
Simulation for a uniform distribution on $[0, 2]$ on the radius : θ_1 constant



$$(\theta_2, \theta_3, \theta_4) = (0, 0.2, 0)$$



$$(\theta_2, \theta_3, \theta_4) = (0, 0, 1)$$



$$(\theta_2, \theta_3, \theta_4) = (-1, -1, 0)$$

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Maximum likelihood method

We observe φ_{Λ_n} , a realization of a **hereditary** marked point process satisfying **[Mod]** in Λ_n .

$$\hat{\theta}_n^{MLE} = \underset{\theta \in \Theta}{\operatorname{argmax}} \frac{1}{Z_{\Lambda_n}(\varphi; \theta)} e^{-V(\varphi_{\Lambda_n}; \theta)}$$

where

$$Z_{\Lambda}(\varphi; \theta) = \sum_{k \geq 0} \frac{1}{k!} \int_{\Lambda} \dots \int_{\Lambda} e^{-V_{\Lambda}(\{x_1, \dots, x_k\}; \theta)} dx_1 \dots dx_k.$$

$Z_{\Lambda_n}(\varphi; \theta)$ is untractable!!

Remarks

- Intensive Monte-Carlo based simulations (Møller(07)) have to be used to estimate $Z(\theta)$.
- Only few theoretical results are available for $\hat{\theta}_n^{MLE}$ (e.g. consistency in general?)

Pseudo-likelihood method (1)

Idea on the lattice (Besag (68), Ripley (88)) : consider the product of the conditional densities in each site conditionally on the other ones.

For point processes JM(94) extended the definition of log-pseudolikelihood function

$$LPL_{\Lambda_n}(\varphi; \theta) = - \int_{\Lambda_n \times \mathbb{M}} e^{-V(x^m | \varphi; \theta)} \mu(dx^m) - \sum_{x^m \in \varphi_{\Lambda_n}} V(x^m | \varphi \setminus x^m; \theta).$$

Define $\hat{\theta}_n^{MPLE} := \underset{\theta \in \Theta}{\operatorname{argmax}} LPL_{\Lambda_n}(\varphi; \theta)$.

Remarks

- “computable” estimate, quick and easy implementation.
- Seems less accurate than the MLE (when available).
- Asymptotic results available for a large class of energies
BCD(08), CD(10), DL(09)

Pseudo-likelihood method (2)

Let Λ_n be a cube with volume growing to $+\infty$.

Proposition BCD(08),CD(10),DL(09)

Under **[Mod]** and **[FR]**, the assumption **[Id-MPLE]**

$$\forall \theta \neq \theta^*, \quad P(V(0^M | \Phi; \theta) \neq V(0^M | \Phi; \theta^*)) > 0,$$

regularity and integrability assumptions on the local energy function

(i) $\hat{\theta}_{MPLE}(\Phi) \xrightarrow{a.s.} \theta^*$.

(ii) if P_{θ^*} is ergodic, $\exists \Sigma_{MPLE} \geq 0 : |\Lambda_n|^{1/2} \left(\hat{\theta}_{MPLE}(\Phi) - \theta^* \right) \xrightarrow{d} \mathcal{N}(0, \Sigma_{MPLE})$.

(iii) If P_{θ^*} is not ergodic but with additional assumptions on Σ_{MPLE} , one can define a consistent estimate of $\hat{\Sigma}_{MPLE}^{-1/2}$ of $\Sigma_{MPLE}^{-1/2}$ and derive a normalized CLT.

Takacs-Fiksel method (1)

Theorem (Georgii-Nguyen-Zessin)

For any $h(\cdot, \cdot; \theta) : \mathbb{S} \times \Omega \rightarrow \mathbb{R}$, for any $\theta \in \Theta$,

$$E_{\theta^*} \left(\sum_{x^m \in \varphi} h(x^m, \varphi \setminus x^m; \theta) \right) = E_{\theta^*} \left(\int_{\mathbb{R}^d \times \mathbb{M}} h(x^m, \varphi; \theta) e^{-V(x^m | \varphi; \theta^*)} \mu(dx^m) \right)$$

Define

$$I_{\Lambda}(\varphi; h, \theta) = \int_{\Lambda \times \mathbb{M}} h(x^m, \varphi; \theta) e^{-V(x^m | \varphi; \theta)} \mu(dx^m) - \sum_{x^m \in \varphi_{\Lambda}} h(x^m, \varphi \setminus x^m; \theta).$$

Idea of TF method : ergodic theorem and GNZ formula $\Rightarrow I_{\Lambda_n}(\varphi; \theta^*) \simeq 0$.

Let us give K test functions $h_k(\cdot, \cdot; \theta) : \mathbb{S} \times \Omega \rightarrow \mathbb{R}$ (for $k = 1, \dots, K$).

$$\hat{\theta}_{TF}(\varphi) := \arg \min_{\theta \in \Theta} \sum_{k=1}^K I_{\Lambda_n}(\varphi; h_k, \theta)^2,$$

Takacs-Fiksel method (2)

$$\hat{\theta}_{TF}(\varphi) := \arg \min_{\theta \in \Theta} \sum_{k=1}^K \left(\int_{\Lambda_n \times \mathbb{M}} h_k(x^m, \varphi; \theta) e^{-V(x^m | \varphi; \theta)} \mu(dx^m) - \sum_{x^m \in \varphi_{\Lambda_n}} h_k(x^m, \varphi \setminus x^m; \theta) \right)^2$$

Remarks and interest of the TF method :

- when $\mathbf{h} = \mathbf{V}^{(1)}$, $\hat{\theta}_{TF}(\varphi) = \hat{\theta}_{MPLE}(\varphi)$.
- quick estimator : for example $h_k(x^m, \varphi; \theta) := \mathbf{1}_{B(0, r_k)}(\|x\|) e^{V(x^m | \varphi; \theta^*)}$.
- allows the identification of the Quermass model with a pertinent choice of test functions allowing to compute the sum term.

Takacs-Fiksel method (2)

$$\hat{\theta}_{TF}(\varphi) := \arg \min_{\theta \in \Theta} \sum_{k=1}^K \left(\int_{\Lambda_n \times \mathbb{M}} h_k(x^m, \varphi; \theta) e^{-V(x^m | \varphi; \theta)} \mu(dx^m) - \sum_{x^m \in \varphi \Lambda_n} h_k(x^m, \varphi \setminus x^m; \theta) \right)^2$$

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- quick estimator : for example $h_k(x^m, \varphi; \theta) := \mathbf{1}_{B(0, r_k)}(\|x\|) e^{V(x^m | \varphi; \theta^*)}$.
- allows the identification of the Quermass model with a pertinent choice of test functions allowing to compute the sum term.
- The consistency and the CLT may be obtained (CDDL(10)) under **[Mod]**, **[FR]**, regularity and integrability assumptions and the identifiability condition **[Id-TF]**

$$\sum_{k=1}^K E \left(h_k(0^M, \Phi; \theta) \left(e^{-V(0^M | \Phi; \theta)} - e^{-V(0^M | \Phi; \theta^*)} \right) \right)^2 = 0 \implies \theta = \theta^*.$$

- Problem with the choice of test functions, **[Id-TF]** may fail! More practical criterion have been proposed in CDDL(10).

Variational principle method (non-marked) Baddeley and Dereudre (10)

VP equation : \forall function $h : \mathbb{S} \times \Omega \rightarrow \mathbb{R}$

$$E_{\theta^*} \left(\sum_{x \in \Phi} \nabla h(x|\varphi \setminus \Phi) \right) = E_{\theta^*} \left(\sum_{x \in \Phi} h(x|\Phi \setminus x) \nabla V(x|\Phi; \theta^*) \right).$$

- If $V(x|\varphi; \theta^*) \rightarrow V(x|\varphi; \theta^*) + c$: same VP equ. \Rightarrow cannot estimate θ_1^* .
- Assume $V(x|\varphi; \theta^*) = \theta_1 + \tilde{\theta}^T \tilde{\mathbf{V}}(x|\varphi)$, VP equ. leads to

$$E_{\theta^*} \left(\sum_{x \in \Phi} \operatorname{div} h(x|\varphi \setminus x) \right) = \tilde{\theta}^T \left(\sum_{x \in \Phi} h(x|\varphi \setminus x) \operatorname{div} \tilde{\mathbf{V}}_i(x|\varphi \setminus x) \right)_{i=1, \dots, p}$$

- Let us give h_1, \dots, h_p and define the vector G and the matrix D by

$$G_k := \sum_{x \in \varphi \wedge_n} \operatorname{div} h_k(x|\varphi \setminus x) \quad \text{and} \quad D_{k,i} := \sum_{x \in \varphi \wedge_n} h_k(x|\varphi \setminus x) \operatorname{div} \tilde{\mathbf{V}}_i(x|\varphi \setminus x)$$

$$\hat{\tilde{\theta}} := D^{-1} G.$$

Rmk : $h_k = \operatorname{div} \tilde{\mathbf{V}}_k \Rightarrow D \geq 0$, and asymptotic results can be obtained.

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Validation through residuals

Baddeley et al (05,08) have proposed to use the GNZ formula as a diagnostic tool :

- 1 Let us give a model and define an estimate $\hat{\theta}$ of θ^* .
- 2 Let us give a test function and define the h -residuals $R_{\Lambda_n}(\varphi; h) = I_{\Lambda_n}(\varphi; h, \hat{\theta})$, i.e.

$$R_{\Lambda_n}(\varphi; h) := \int_{\Lambda_n \times \mathbb{M}} h(x^m, \varphi; \hat{\theta}) e^{-V(x^m | \varphi; \hat{\theta})} \mu(dx^m) - \sum_{x^m \in \varphi_{\Lambda_n}} h(x^m, \varphi \setminus x^m; \hat{\theta})$$

If the model is valid, then one may expect that $R_{\Lambda_n}(\varphi; h)/|\Lambda_n| \simeq 0$!

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Examples :

- Classical examples $h = 1, e^V, e^{V/2}$ leading to the raw, the inverse and the Pearson residuals.
- A more evolved one : let $h_r(x^m, \varphi; \theta) := \mathbf{1}_{[0,r]}(d(x^m, \varphi)) e^{V(x^m | \varphi; \theta)}$ where $d(x^m, \varphi) = \inf_{y^m \in \varphi} \|y - x\|$, $R_{\Lambda_n}(\varphi; h_r)$ corresponds to a difference of two estimates of the empty space function $F(r) := P(d(0^M, \Phi) \leq r)$

Measures of departures to the true model

- ① Objective 1 : fix some test function h (ex : $h = 1$) and split $\Lambda_n = \cup_{j \in \mathcal{J}} \Lambda_j$

$\hat{\theta}$ is based on ϕ_{Λ_n} and $\Lambda_n = \cup_{j \in \mathcal{J}} \Lambda_{j,n}$

		$R_{\Lambda_{1,n}}(\phi, h, \hat{\theta})$
	$R_{\Lambda_{2,n}}(\phi, h, \hat{\theta})$	
$R_{\Lambda_{s,n}}(\phi, h, \hat{\theta})$		

- ① compute the h -residuals in each Λ_j

- ② construct $\|\mathbf{R}_1\|^2$ where

$$\mathbf{R}_1 := \left(R_{\Lambda_j}(\varphi, h, \hat{\theta}_n) \right)_{j \in \mathcal{J}}.$$

- ② Objective 2 : fix $\mathbf{h} = (h_1, \dots, h_s)^T$ (ex : $h_j = h_{r_j}$, $r_1 < r_2 < \dots < r_s$)

$\hat{\theta}$ is based on φ_{Λ_n}

$$\begin{cases} R_{\Lambda_n}(\phi, h_1, \hat{\theta}_n) \\ \vdots \\ R_{\Lambda_n}(\phi, h_j, \hat{\theta}_n) \\ \vdots \\ R_{\Lambda_n}(\phi, h_s, \hat{\theta}_n) \end{cases}$$

- ① compute the s h_j -residuals in Λ_n

- ② construct $\|\mathbf{R}_2\|^2$ where

$$\mathbf{R}_2 := \left(R_{\Lambda_n}(\varphi, h_j, \hat{\theta}_n) \right)_{j=1, \dots, s}.$$

Towards goodness-of-fit tests

Under similar assumptions as previously and with general assumptions on $\hat{\theta}$ (essentially consistency and CLT)

Proposition (CL(10))

- (i) As $n \rightarrow +\infty$, $\mathbf{R}_1(\Phi; h)/|\Lambda_{0,n}|$ and $\mathbf{R}_2(\Phi; \mathbf{h})/|\Lambda_n|$ converge a.s. to 0.
(ii) If P_{θ^*} is not ergodic (but with additional assumptions) a normalized CLT holds for \mathbf{R}_1 and \mathbf{R}_2 , leading to

$$|\Lambda_{0,n}|^{-1} \|\widehat{\Sigma}_1^{-1/2} \mathbf{R}_1(\Phi; h)\|^2 \xrightarrow{d} \chi_{|\mathcal{J}|}^2 \quad \text{and} \quad |\Lambda_n|^{-1} \|\widehat{\Sigma}_2^{-1/2} \mathbf{R}_2(\Phi; \mathbf{h})\|^2 \xrightarrow{d} \chi_S^2.$$

- $\Sigma_1 = \lambda_{Inn} \mathbf{I}_{|\mathcal{J}|} + |\mathcal{J}|^{-1} (\lambda_{Res} - \lambda_{Inn}) \mathbf{J}$ with $\mathbf{J} = \mathbf{e}\mathbf{e}^T$ and $\mathbf{e} = (1, \dots, 1)^T$.
- $\lambda_{Inn} = f(P_{\theta^*}, V, h)$ and $\lambda_{Res} = f(\theta^*, V, h, \hat{\theta})$.
- This form suggested us to study the centered residuals for which we may prove

$$|\Lambda_{0,n}|^{-1} \widehat{\lambda}_{Inn}^{-1} \|\mathbf{R}_1(\Phi; h) - \overline{\mathbf{R}}_1(\Phi; h)\|^2 \xrightarrow{d} \chi_{|\mathcal{J}|-1}^2.$$