Kernel density estimation in adaptive tracking

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Universities of Bordeaux and Rouen, France 47th IEEE Conference on Decision and Control Cancun, Mexico, December 10, 2008

Outline



Kernel density estimation



Estimation and adaptive control

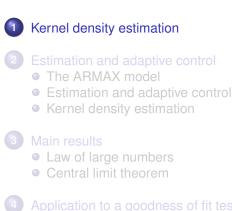
- The ARMAX model
- Estimation and adaptive control
- Kernel density estimation

Main results

- Law of large numbers
- Central limit theorem

Application to a goodness of fit test

Outline



Goal

Let (X_n) be a sequence of **iid** random variables with **unknown** density function *f*.

Goal

Estimate the density function f by a kernel density estimator.

Let *K* be a **nonnegative**, **bounded**, **Lipschitz** function called **Kernel**, such that

$$\int_{\mathbb{R}} \mathcal{K}(x) \, dx = 1, \qquad \int_{\mathbb{R}} x \mathcal{K}(x) \, dx = 0,$$
$$\int_{\mathbb{R}} \mathcal{K}^{2}(x) \, dx = \tau^{2}.$$

Choices for the kernel

• Uniform kernel

$$\mathcal{K}_a(x) = \begin{cases} \frac{1}{2a} & \text{if } |x| \leq a, \\ 0 & \text{otherwise.} \end{cases}$$

• Epanechnikov kernel

$$\mathcal{K}_b(x) = \begin{cases} \frac{3}{4b} \left(1 - \frac{x^2}{b^2} \right) & \text{if } |x| \leq b, \\ 0 & \text{otherwise} \end{cases}$$

• Gaussian kernel

$$K_c(x) = \frac{1}{c\sqrt{2\pi}} \exp\left(-\frac{x^2}{2c^2}\right).$$

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Parzen-Rosenblatt or Wolverton-Wagner

Let (h_n) be a sequence of positive real numbers decreasing to zero called **bandwidth**. We can estimate the density *f* by the **Parzen-Rosenblatt** estimator given for all $x \in \mathbb{R}$ by

$$\widetilde{f}_n(x) = \frac{1}{nh_h} \sum_{i=1}^n K\Big(\frac{X_i - x}{h_n}\Big).$$

We can also estimate the density *f* by the Wolverton-Wagner estimator given for all $x \in \mathbb{R}$ by

$$\widehat{f}_n(\mathbf{x}) = \frac{1}{n} \sum_{i=1}^n \frac{1}{h_i} K\left(\frac{\mathbf{X}_i - \mathbf{x}}{h_i}\right).$$

On Wolverton-Wagner

Theorem

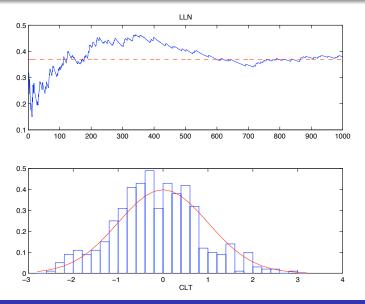
Assume that f is derivable with bounded derivative. If the bandwidth $h_n = 1/n^{\alpha}$ with $0 < \alpha < 1$, we have

(LLN)
$$\lim_{n\to\infty}\widehat{f}_n(x)=f(x) \quad a.s.$$

In addition, if $1/5 < \alpha < 1$, we also have

CLT)
$$\sqrt{nh_n}(\widehat{f}_n(x) - f(x)) \xrightarrow{\mathcal{L}} \mathcal{N}\left(0, \frac{\tau^2 f(x)}{1 + \alpha}\right)$$

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The ARMAX model Estimation and adaptive control Kernel density estimation

Outline



Kernel density estimation



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The ARMAX model Estimation and adaptive control Kernel density estimation

Consider the *d*-dimensional **ARMAX(p,q,r)** model given by

 $\boldsymbol{A}(\boldsymbol{R})\boldsymbol{X}_n = \boldsymbol{B}(\boldsymbol{R})\boldsymbol{U}_n + \boldsymbol{C}(\boldsymbol{R})\boldsymbol{\varepsilon}_n$

where *R* is the shift-back operator, X_n is the system output, U_n is the system input and ε_n is the driven noise,

- $A(R) = I_d A_1 R \cdots A_p R^p$,
- $B(R) = B_1 R + B_2 R^2 + \dots + B_a R^q$,
- $C(R) = I_d C_1 R \cdots C_r R^r$

where A_i , B_j , and C_k are unknown matrices. We assume that the **high frequency gain** B_1 is known with $B_1 = I_d$.

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$$A(R) = I_d - A_1 R - \cdots - A_p R^p$$
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where A_i , B_j , and C_k are unknown matrices. We assume that the **high frequency gain** B_1 is known with $B_1 = I_d$.

The ARMAX model Estimation and adaptive control Kernel density estimation

The unknown parameter of the model is given by

$$\theta^t = (A_1, \ldots, A_p, B_2, \ldots, B_q, C_1, \ldots, C_r).$$

The **ARMAX(p,q,r)** model can be rewritten as

 $X_{n+1} = \theta^t \Psi_n + U_n + \varepsilon_{n+1},$

where $\Psi_n = (X_n^p, U_n^q, \varepsilon_n^r)^t$ with

$$\begin{aligned} X_n^p &= (X_n^t, \dots, X_{n-p+1}^t), \\ U_n^q &= (U_{n-1}^t, \dots, U_{n-q+1}^t), \\ \varepsilon_n^r &= (\varepsilon_n^t, \dots, \varepsilon_{n-r+1}^t). \end{aligned}$$

The ARMAX model Estimation and adaptive control Kernel density estimation

Causality and Passivity

Definition

The matrix polynomial *B* is **causal** if for all $z \in \mathbb{C}$ with $|z| \leq 1$

 $\det(z^{-1}B(z))\neq 0.$

Definition

The matrix polynomial *C* is **passif** if for all $z \in \mathbb{C}$ with |z| = 1

 $\det(C(z)) \neq 0$ and $C^{-1}(z) > \frac{1}{2}I_d$

The ARMAX model Estimation and adaptive control Kernel density estimation

Extended least squares

We estimate θ by the **extended least squares** estimator

$$\widehat{\theta}_{n+1} = \widehat{\theta}_n + S_n^{-1} \Phi_n (X_{n+1} - U_n - \widehat{\theta}_n^t \Phi_n)^t,$$

$$\widehat{\varepsilon}_{n+1} = X_{n+1} - U_n - \widehat{\theta}_n^t \Phi_n,$$

where the vector $\Phi_n = (X_n^p, U_n^q, \hat{\varepsilon}_n^r)^t$ with $\hat{\varepsilon}_n^r = (\hat{\varepsilon}_n^t, \dots, \hat{\varepsilon}_{n-r+1}^t)$,

$$S_n = \sum_{i=0}^n \Phi_i \Phi_i^t + S,$$

where S is a positive definite and deterministic matrix.

The ARMAX model Estimation and adaptive control Kernel density estimation

Adaptive Control

The role played by U_n is to force X_n to track step by step a given trajectory (x_n) . We make use of the **adaptive tracking control**

 $\boldsymbol{U}_n = \boldsymbol{x}_{n+1} - \widehat{\theta}_n^t \boldsymbol{\Phi}_n.$

Then, the closed-loop system is given by

 $X_{n+1} - X_{n+1} = \pi_n + \varepsilon_{n+1}$

where the prediction error

$$\pi_n = (\theta - \widehat{\theta}_n)^t \Phi_n.$$

The ARMAX model Estimation and adaptive control Kernel density estimation

We assume that (ε_n) is a sequence of **iid** random vectors with **unknown density** *f*. If (ε_n) were observable, we could estimate *f* by

$$f_n(\mathbf{x}) = \frac{1}{n} \sum_{i=1}^n \frac{1}{h_i^d} K\left(\frac{\varepsilon_i - \mathbf{x}}{h_i}\right).$$

However, ε_{n+1} is unobservable but it can estimated by

$$\widehat{\varepsilon}_{n+1} = X_{n+1} - U_n - \widehat{\theta}_n^t \Phi_n = X_{n+1} - X_{n+1}.$$

Consequently, we can use the Wolverton-Wagner estimator

$$\widehat{f}_n(\mathbf{x}) = \frac{1}{n} \sum_{i=1}^n \frac{1}{h_i^d} K\left(\frac{X_i - X_i - \mathbf{x}}{h_i}\right).$$

Law of large numbers Central limit theorem

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- Estimation and adaptive control
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- Main results
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Application to a goodness of fit test

Law of large numbers Central limit theorem

Uniform law of large numbers

Theorem

Assume that *f* is positive and differentiable with **bounded** gradient and that (ε_n) has finite moment of order > 2. If the bandwidth $h_n = 1/n^{\alpha}$ with $\alpha \in]0, 1/d[$, then

(LLN)

$$\lim_{n\to\infty}\sup_{x\in\mathbb{R}^d}|\widehat{f}_n(x)-f(x)|=0 \quad a.s.$$

Law of large numbers Central limit theorem

Central limit theorem

Theorem

Assume that *f* is positive and differentiable with **bounded** gradient and that (ε_n) has finite moment of order > 2. If the bandwidth $h_n = 1/n^{\alpha}$ with $\alpha \in [1/(d+2), 1/d[$, then

$$G_n(x) = \sqrt{nh_n^d}(\widehat{f}_n(x) - f(x)) \xrightarrow{\mathcal{L}} \mathcal{N}\left(0, \frac{\tau^2 f(x)}{1 + \alpha d}\right) = G(x).$$

In addition, for N distinct points x_1, \ldots, x_N of \mathbb{R}^d , we also have

(MCLT)
$$(G_n(x_1), \cdots, G_n(x_N)) \xrightarrow{\mathcal{L}} (G(x_1), \cdots, G(x_N))$$

where $G(x_1), \ldots, G(x_N)$ are independent.

Outline



goodness of fit test

We wish to test

 \mathcal{H}_0 : $\langle f = f_0 \rangle$ versus \mathcal{H}_1 : $\langle f \neq f_0 \rangle$

where f_0 is a given density function. Our statistical test is

$$T_n(N) = \frac{1}{\tau^2 \ell_h} \sum_{i=1}^N \frac{(\hat{f}_n(x_i) - f_0(x_i))^2}{\hat{f}_n(x_i)}$$

where x_1, \ldots, x_N are *N* distinct points of \mathbb{R}^d and

$$\ell_h = \frac{1}{1 + \alpha d}.$$

Theorem

Assume that f is positive and differentiable with **bounded** gradient and that (ε_n) has finite moment of order > 2. If the bandwidth $h_n = 1/n^{\alpha}$ with $\alpha \in]1/(d+2), 1/d[$, then under \mathcal{H}_0

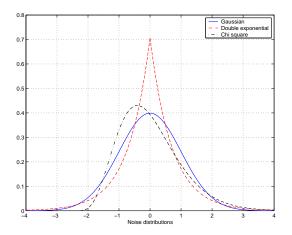
$$nh_n^d T_n(N) \xrightarrow{\mathcal{L}} \chi^2(N).$$

In addition, under \mathcal{H}_1 and if one can find some point x of \mathbb{R}^d in $\{x_1, x_2, \ldots, x_N\}$ such that $f(x) \neq f_0(x)$, then $T_n(N) \rightarrow \sigma^2$ a.s.

$$\sqrt{nh_n^d}(T_n(N) - \sigma^2) \stackrel{\mathcal{L}}{\longrightarrow} \mathcal{N}(\mathbf{0}, \lambda^2)$$

$$\sigma^{2} = \frac{1}{\tau^{2}\ell_{h}} \sum_{i=1}^{N} \frac{(f(x_{i}) - f_{0}(x_{i}))^{2}}{f(x_{i})}, \quad \lambda^{2} = \frac{1}{\tau^{2}\ell_{h}} \sum_{i=1}^{N} \frac{(f^{2}(x_{i}) - f_{0}^{2}(x_{i}))^{2}}{f^{3}(x_{i})}.$$

Simulations



Noise distributions

• Gaussian

$$f_0(x) = rac{1}{\sqrt{2\pi}} \exp\Bigl(-rac{x^2}{2}\Bigr),$$

Double exponential

$$f_1(x) = \frac{1}{\sqrt{2}} \exp\left(-\sqrt{2}|x|\right),$$

• Chi square

$$f_2(x) = \begin{cases} \frac{9}{5}(x+\sqrt{6})^5 \exp\left(-\sqrt{6}(x+\sqrt{6})\right) & \text{if } |x| \ge -\sqrt{6}, \\ 0 & \text{otherwise.} \end{cases}$$

Noise distributions

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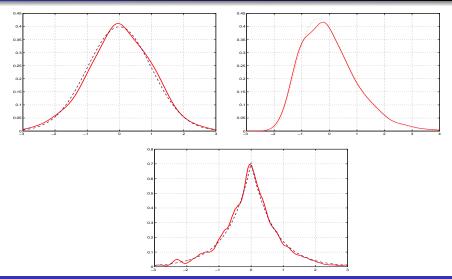
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Law of large numbers



Bercu and Portier

Kernel density estimation in adaptive tracking

ARX Goodness of fit test

$X_{n+1} = \theta X_n + U_n + \varepsilon_{n+1}$

Table: Results under \mathcal{H}_0 and \mathcal{H}_1 with test level 5%.

| | <i>n</i> = 200, <i>N</i> = 8 | | | | <i>n</i> = 1000, <i>N</i> = 22 | |
|------------------|------------------------------|------------------|------------------|------------------|--------------------------------|------------------|
| | $\mathcal{H}f_0$ | $\mathcal{H}f_1$ | $\mathcal{H}f_2$ | $\mathcal{H}f_0$ | $\mathcal{H}f_1$ | $\mathcal{H}f_2$ |
| G f ₀ | 3.8% | 35.7% | 28% | 3.7% | 99.7% | 98.2% |
| $\mathcal{G}f_1$ | 45.8% | 5.5% | 71.5% | 100% | 5% | 100% |
| $\mathcal{G}f_2$ | 21.2% | 54.5% | 3.2% | 96.7% | 100% | 5.1% |

NARX Goodness of fit test

$$X_{n+1} = \theta X_n^2 + U_n + \varepsilon_{n+1}$$

Table: Results under \mathcal{H}_0 and \mathcal{H}_1 with test level 5%.

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| | $\mathcal{H}f_0$ | $\mathcal{H}f_1$ | $\mathcal{H}f_2$ | $\mathcal{H}f_0$ | $\mathcal{H}f_1$ | $\mathcal{H}f_2$ |
| Gf ₀ | 3% | 37.1% | 28.5% | 4.3% | 99.5% | 98.6% |
| $\mathcal{G}f_1$ | 44.6% | 5.2% | 72% | 100% | 5.1% | 100% |
| $\mathcal{G}f_2$ | 19.8% | 58.3% | 3.7% | 97.2% | 100% | 5% |