

# A Robust Maximal $F$ -ratio Statistic to Detect Clusters Structure

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# Outline

- 1 Introduction
- 2 The Maximal  $F$ -ratio
- 3 A Trimmed version of the Maximal  $F$ -ratio
- 4 Asymptotics for the Trimmed Maximal  $F$ -ratio
- 5 Hypothesis Testing
- 6 An Example in Genetics

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# F-ratio statistics

- F-ratio is a widely used tool in Statistical Data Analysis

$$F = \frac{\text{Between Groups Variance}}{\text{Within Groups Variance}}$$

$$F = \frac{B_k}{W_k} = \frac{\sum_{j=1}^k n_j \|\bar{y}_j - \bar{y}\|^2}{\sum_{j=1}^k \sum_{i=1}^{n_j} \|y_{i,j} - \bar{y}_j\|^2}$$

- **Supervised** (known group ownerships):
  - **ANOVA**  
Testing the existence of differences between the groups means.
  - **Linear Discriminant Analysis**  
Derivation of the canonical variates.
- **Unsupervised** (unknown group ownerships):
  - Maximal F-Ratio in **Cluster Analysis**  
Testing the number of clusters.

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$k$ -means (I)

- McQueen (1967), Hartigan and Wong (1979).
- Well-known and widely used (non-hierarchical) clustering method.
- $x_1, x_2, \dots, x_n$ :  $p$ -variate observations.
- $C_1, \dots, C_k$ : partition into  $k$  groups.  $m_1, \dots, m_k$ : groups sample means.
- $k$ -means  $m_1^n, \dots, m_k^n$ : Solution of the **Minimization** problem

$$\min_{C_1, \dots, C_k} \sum_{j=1}^k \sum_{x_i \in C_j} \|x_i - m_j\|^2 = \min_{m_1, \dots, m_k} \sum_{i=1}^n \inf_{1 \leq j \leq k} \|x_i - m_j\|^2$$

- **Within-groups** sum of squares:

$$W_k^n = \sum_{i=1}^n \inf_{1 \leq j \leq k} \|x_i - m_j^n\|^2$$

## $k$ -means (II)

- $x_1, x_2, \dots, x_n$ :  $p$ -variate observations
- $C_1, \dots, C_k$ : partition into  $k$  groups.  $m_1, \dots, m_k$ : groups sample means.
- $m^n$ : overall mean
- $n_1, \dots, n_k$ : groups sizes
- $k$ -means  $m_1^n, \dots, m_k^n$ : Solution of **Maximization** problem

$$\max_{C_1, \dots, C_k} \sum_{j=1}^k n_j \|m_j - m^n\|^2$$

- **Between-groups** sum of squares:

$$B_k^n = \sum_{j=1}^k n_j \|m_j^n - m^n\|^2$$

# Maximal $F$ -Ratio Statistic (I)

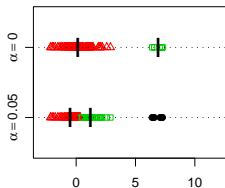
$$R_k^n = \frac{B_k^n}{W_k^n}$$

- Maximal  $F$ -Ratio Statistic to **test two clusters structure**:
  - Engelman and Hartigan (1969) proposed a **test** for “clusters structure” for univariate data and  $k = 2$ .
  - The test divides the sample into two subsets **maximizing the likelihood ratio** that the two subsets are sampled from two normals with different means, against the null hypothesis that the means are equal.
  - The test **reject  $H_0$**  when the 2-means centers are “**separated**” enough.
- Maximal  $F$ -Ratio Statistic to **check  $k$  clusters structure**:
  - Calinski and Harabasz (1974): Extension to check “ $k - 1$  or less groups” against “at least  $k$  groups”.
    - It is **not straightforward**: We need to specify the distribution in  $H_0$ .
    - “Pseudo  $F$ -test”: **Descriptive Criterion** for choosing  $k$ .
    - **Good results in Milligan and Cooper (1985)’s study** among 30 procedures for determining the number of clusters.
  - Hartigan (1978) proved the **asymptotic normality** for the  $k$ -mean centers and the Maximal  $F$ -ratio in the univariate case.

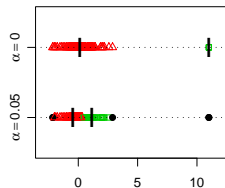


# Maximal $F$ -Ratio Statistic (II)

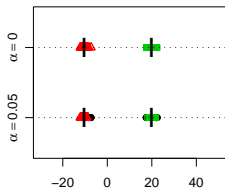
- Extension to the **multivariate** setting
  - Lee (1979) extended the Maximal  $F$ -Ratio principle to test clusters structure in the multivariate case.
  - Pollard (1982) extended Hartigan's results to the multivariate case by using Empirical Process Theory.
- Main **drawbacks**:
  - Hartigan and Pollard's results need some **moment conditions** on the underlying distribution to be applied.
  - **Lack of Robustness**:  $k$ -means and Maximal  $F$ -ratio can be severely affected by a small amount of anomalous observations



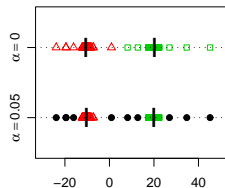
(a)



(b)



(c)



(d)

(a) 100 observations: 95%  $N(0, 1)$  and 5% outliers.

(b) 100 observations: 99%  $N(0, 1)$  and 1% remote outliers.

(c) 200 observations drawn from a normal mixture.

(d) 200 observations: 95% from a normal mixture and 5% background noise.

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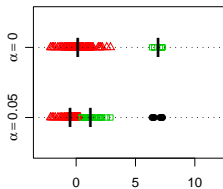
# Robust Maximal F-ratio based on Trimmed $k$ -means

- **Trimming** data in **clustering** problems is **not an easy task**.
- Gordaliza (1991) introduced “**impartial**” **trimming**.
- Cuesta-Albertos et al (1997) extended the idea to clustering.
- **Trimmed  $k$ -means**:  $m_1^n, \dots, m_k^n$  solution of the Minimization problem:

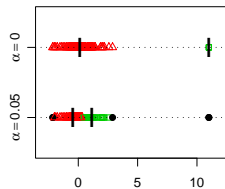
$$\min_{\mathbf{Y}} \min_{m_1, \dots, m_k} \sum_{x_i \in \mathbf{Y}} \inf_{1 \leq j \leq k} \|x_i - m_j\|^2,$$

- $\alpha \in (0, 1)$  is the trimming proportion.
  - $\mathbf{Y}$  ranges over all subsets  $\mathbf{Y} \subset \{x_1, \dots, x_n\}$  with  $[n(1 - \alpha)]$  elements.
  - The trimmed  $k$ -means  $m_1^n, \dots, m_k^n$  induce a partition of the non-trimmed observations onto  $k$  clusters  $C_1^n \cup \dots \cup C_k^n$ .
  - The clusters are balls having the same radius.
- **Trimmed Maximal F-Ratio**:

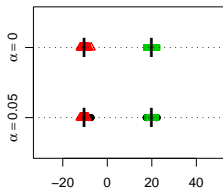
$$R_k^n(\alpha) := \frac{B_k^n(\alpha)}{W_k^n(\alpha)} := \frac{\sum_{j=1}^k n_j \left\| m_j^n - m^n \right\|^2}{\sum_{j=1}^k \sum_{x_i \in C_j^n} \left\| x_i - m_j^n \right\|^2}$$



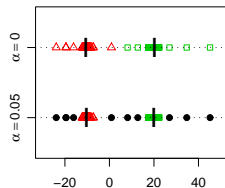
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# Population version

- Population version of Trimmed  $k$ -means.  $m_1^0, \dots, m_k^0$  solution of:

$$A_0 : \min_{P(A_0)=1-\alpha} \min_{m_1^0, \dots, m_k^0} \int_{A_0} \inf_{1 \leq j \leq k} \|x - m_j^0\|^2 dP(x)$$

- Population version of Trimmed Maximal  $F$ -ratio:

$$R_k^0(\alpha) := \frac{B_k^0(\alpha)}{W_k^0(\alpha)} := \frac{\sum_{j=1}^k P(C_j^0) \|m_j^0 - m^0\|^2}{\sum_{j=1}^k \int_{C_j^0} \|x - m_j^0\|^2 dP(x)}.$$

- $C_1^0, \dots, C_k^0$ : partition of the non trimmed area into  $k$  disjoint groups.
- $m_1^0, \dots, m_k^0$ : groups sample means (Trimmed  $k$ -means).
- $m^0$ : overall trimmed mean.

# Asymptotic properties

- **Theorem: Consistency**

Let  $\alpha \in (0, 1)$  and  $P$  be a continuous probability distribution and assume that there exists a unique  $\alpha$ -trimmed  $k$ -mean. Then

$$R_k^n(\alpha) \rightarrow R_k^0(\alpha) \text{ } P\text{-a.e.}$$

- **Theorem: Asymptotic Normality**

If  $P$  has bounded density not identically null on the boundary of the optimal trimming set and there exists a unique  $\alpha$ -trimmed  $k$ -mean, then:

$$\sqrt{n}(R_k^n(\alpha) - R_k^0(\alpha)) \xrightarrow{\mathcal{L}} N(0, LVL')$$

- Proof based on Empirical Processes Theory.
- No moment conditions on  $P$  are needed.



Asymptotic means and variances of  $R_k^n(\alpha)$  $k = 2$ 

$\alpha$ :	.0001	.001	.01	.05	.1	.2
Asymp. mean	1.756	1.776	1.877	2.097	2.263	2.483
Asymp. var.	6.515	6.485	7.605	12.841	19.462	34.475

 $k = 3$ 

$\alpha$ :	.0001	.001	.01	.05	.1	.2
Asymp. mean	4.273	4.347	4.698	5.412	5.921	6.568
Asymp. var.	74.775	75.119	79.328	95.600	115.570	160.341

**Table:** Asymptotic means and variances of the maximal  $F$ -ratio for the  $N(0, 1)$  distribution,  $k = 2$  and  $k = 3$  and different trimming levels.

Hartigan (1978)'s values for the untrimmed ( $\alpha = 0$ ) case, when  $k = 2$ :

- Asymptotic mean:  $2/(\pi - 2) = 1.752$

- Asymptotic variance:  $8/\pi \cdot (1 - 3/\pi)/(1 - 2/\pi)^4 = 6.582$

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# Hypothesis Testing

- Asymptotic results can be used to test  $H_0 : k = 1$  against  $H_1 : k > 1$ .
- Reject  $H_0$  when  $R_k^n(\alpha)$  is greater than the (asymptotic) critical value.
- The test depends on the assumed model for  $H_0$ .
- Large sample size  $n$  will be required because of the asymptotic character of the results.
- We conduct a simulation study in the univariate normal case to analyze:
  - The empirical power of the test.
  - The gain in robustness provided by the trimming.
  - The behavior for finite sample sizes.

## Empirical powers

$\alpha:$		$\varepsilon = .0$			$\varepsilon = .01$		
		0	.01	.05	0	.01	.05
$D = 0$	$n = 500$	.08	.08	.05	.87	.33	.02
	$n = 1000$	.06	.06	.04	.95	.26	.01
	$n = 10000$	.05	.05	.04	1.00	.03	.00
$D = 2$	$n = 500$	.94	.90	.76	.41	.60	.66
	$n = 1000$	1.00	.99	.94	.42	.62	.89
	$n = 10000$	1.00	1.00	1.00	.47	.63	1.00
$D = 3$	$n = 500$	1.00	1.00	1.00	.06	.67	1.00
	$n = 1000$	1.00	1.00	1.00	.01	.72	1.00
	$n = 10000$	1.00	1.00	1.00	.00	.98	1.00

**Table:** Empirical powers based on 5000 random samples.

**Significance level: 0.05.**

$$f_{D,\varepsilon}(x) = (1 - \varepsilon)f_D(x) + \varepsilon\varphi(x - 20)$$

$$f_D(x) = 0.5\varphi(x) + 0.5\varphi(x - D) \text{ with } \varphi(\cdot) \text{ being the density of } N(0, 1).$$

Powers obtained using the asymptotic distributions of  $R_2^n$  and  $R_2^n(\alpha)$ .

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# Ordering and selecting genes

- Characteristics of **genetical studies**:
  - Genetical studies usually involve a **large number of genes**.
  - The **presence of noise** is frequent (genomic and proteomic data are frequently affected by measurement errors)
- **Selection of genes**:
  - We **look for informative genes** exhibiting differences between the expression levels across the individuals in the study.
  - When group ownerships are known, **ordering genes according to their  $F$ -Ratios** can be used as a first screening phase in gene selection (Dudoit et al. 2002).
  - When group ownerships are unknown, we can resort to **Maximal  $F$ -Ratios** to **order** genes and **select** a reduced set of informative genes to apply clustering methods.
  - The presence of **noise** suggests the use of **Robust Maximal  $F$ -Ratios** to order and select genes.

# Golub's leukemia dataset (I)

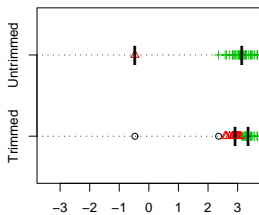
- Example based on the [training set](#) of the famous [Golub's leukemia dataset](#) (Golub et al 1999).
- We study the data subset *leukemia* which accompanies the contributed package [supclust](#) in the [R-project repository](#). (Dettling and Bühlmann 2004)
- Study of gene expression in [two types](#) of acute leukemias:
  - Acute lymphoblastic leukemia ([ALL](#)).
  - Acute myeloid leukemia ([AML](#)).
- The data set consists of [250](#) gene expressions measured on [38 individuals](#).
- We will assume that [no leukemia type classification is available](#) (unsupervised study).

## Golub's leukemia dataset (II)

- We obtain an “importance” index of genes according to the values of the statistics
  - Maximal  $F$ -Ratio Statistic:  $R_2^n$  (Rank)
  - Trimmed Maximal  $F$ -Ratio Statistic:  $R_2^n(.08)$  (T-Rank).
- The trimming size  $\alpha = .08$  is fixed to avoid the influence of the 3 most outlying observations.
- Genes with smallest values in Rank and T-Rank are marked as the most “interesting” ones.

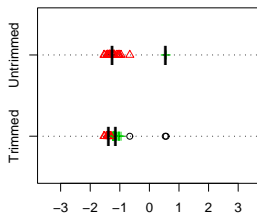


Gene 1



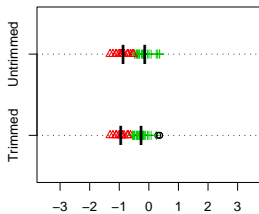
Rank: 48 and T-Rank: 216

Gene 152



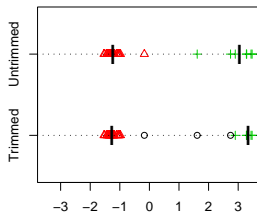
Rank: 17 and T-Rank: 243

Gene 172



Rank: 130 and T-Rank: 137

Gene 230



Rank: 1 and T-Rank: 1

## Ordering subsets of $p$ genes

- The statistic  $R_2^n(\alpha)$  can be computed from  $p$ -dimensional samples.
- The trimmed maximal  $F$ -ratio can be used for ranking subsets of  $p$  variables which jointly serve to detect interesting clusters structures.
- This is a computationally hard problem if the number of genes  $J$  is large as long as  $\binom{J}{p}$  subsets need to be explored.

García-Escudero, L. A., Gordaliza A., Mayo-Iscar, A. and Matrán, C. (2009). A Robust Maximal  $F$ -ratio Statistic to Detect Clusters Structure. *Communications in Statistics - Theory and Methods*, **38**, 682-694.

Thank you very much

Merci beaucoup